Problem set 3

• Problem 1 (4 pts)

It was mentioned in the lecture that a cubic regression spline with two knots at $\xi_1 < \xi_2$ can be represented as

$$f(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \beta_4 (x - \xi_1)_+^3 + \beta_5 (x - \xi_2)_+^3$$

where

$$(x - \xi_k)_+^3 = \begin{cases} (x - \xi_k)^3 & \text{if } x > \xi_k \\ 0 & \text{otherwise} \end{cases}, \quad \forall \ k \in \{1, 2\}.$$
 (0.1)

We now verify f(x) is indeed a cubic regression spline by proving:

- 1. (1 pt) f(x) can be written as three cubic polynomials for $x \leq \xi_1$, $\xi_1 < x \leq \xi_2$ and $x > \xi_2$.
- 2. (1 pt) the three polynomials have the same value at each knot.
- 3. (1 pt) the three polynomials have the same first order derivative at each knot.
- 4. (1 pt) the three polynomials have the same second order derivative at each knot.

SOLUTION:

1. According to the definition of f(x), we have

$$f(x) = \begin{cases} f_1(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 & \text{if } x \le \xi_1 \\ f_2(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \beta_4 (x - \xi_1)^3 & \text{if } \xi_1 < x \le \xi_2 \\ f_3(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \beta_4 (x - \xi_1)^3 + \beta_5 (x - \xi_2)^3 & \text{if } x > \xi_2 \end{cases}$$

as desired.

- 2. This follows immediately from the previous part by letting $x = \xi_1$ and $x = \xi_2$.
- 3. Since

$$\frac{df_1(x)}{dx} = \beta_1 + 2\beta_2 x + 3\beta_3 x^2,
\frac{df_2(x)}{dx} = \beta_1 x + 2\beta_2 x + 3\beta_3 x^2 + 3\beta_4 (x - \xi_1)^2,
\frac{df_3(x)}{dx} = \beta_1 x + 2\beta_2 x + 3\beta_3 x^2 + 3\beta_4 (x - \xi_1)^2 + 3\beta_5 (x - \xi_2)^2,$$

 $f_1'(\xi_1) = f_2'(\xi_1)$ and $f_2'(\xi_2) = f_3'(\xi_2)$ follow by taking $x = \xi_1$ and $x = \xi_2$.

4. Since

$$\frac{d^2 f_1(x)}{dx^2} = 2\beta_2 + 6\beta_3 x,$$

$$\frac{d^2 f_2(x)}{dx^2} = 2\beta_2 + 6\beta_3 x + 6\beta_4 (x - \xi_1)$$

$$\frac{d^2 f_2(x)}{dx^2} = 2\beta_2 + 6\beta_3 x + 6\beta_4 (x - \xi_1) + 6\beta_5 (x - \xi_2),$$

taking $x = \xi_1$ and $x = \xi_2$ leads to $f_1''(\xi_1) = f_2''(\xi_1)$ and $f_2''(\xi_2) = f_3''(\xi_2)$.

• Problem 2 (7 pts)

Consider the classification problem with the label of Y belong to $\mathcal{C} := \{1, 2, ..., K\}$ and any realization x of $X \in \mathbb{R}^p$. The Bayes classifier f^* at X = x is defined as

$$f^*(x) := \underset{f(x) \in \mathcal{C}}{\operatorname{argmin}} \ \mathbb{E} \Big[1\{Y \neq f(X)\} \mid X = x \Big].$$

1. (2 pts) Prove that

$$f^*(x) = \underset{k \in \mathcal{C}}{\operatorname{argmax}} \ \mathbb{P}(Y = k \mid X = x).$$

2. (1 pt) Prove that the Bayes error at X = x equals to

$$1 - \max_{k \in \mathcal{C}} \ \mathbb{P}(Y = k \mid X = x).$$

3. (2 pts) Consider that K=3. For a fixed x_0 , assume that

$$\mathbb{P}(Y = 1 \mid X = x_0) = 0.5$$

$$\mathbb{P}(Y = 2 \mid X = x_0) = 0.3$$

$$\mathbb{P}(Y = 3 \mid X = x_0) = 0.2.$$

State the Bayes classifier at $X = x_0$ and compute its error at $X = x_0$.

4. (2 pts) Consider a naive classifier \hat{f} , called random guessing, which randomly picks one label from $\mathcal{C} = \{1, 2, 3\}$ with equal probability. Compute its expected error rate at $X = x_0$ and compare it with the Bayes error at $X = x_0$ in part 3.

SOLUTION:

1. **Proof.** We have

$$\begin{split} \mathbb{P}(Y = f(X) \mid X = x) &= \sum_{k \in \mathcal{C}} \mathbb{P}(Y = f(X) = k \mid X = x) \\ &= \sum_{k \in \mathcal{C}} \mathbb{P}(Y = k \mid X = x) \ 1\{f(x) = k\} \\ &\leq \max_{k \in \mathcal{C}} \mathbb{P}(Y = k \mid X = x). \end{split}$$

Moreover, the maximal value is achieved when $f(x) = k^*$ for

$$k^* = \operatorname*{argmax}_{k \in \mathcal{C}} \mathbb{P}(Y = k \mid X = x).$$

This proves the first claim.

2. **Proof.** Regarding the second claim, by similar reasoning,

$$\mathbb{P}(Y \neq f^*(X) \mid X = x) = 1 - \mathbb{P}(Y = f^*(X) \mid X = x)$$

$$= 1 - \sum_{k \in \mathcal{C}} \mathbb{P}(Y = k \mid X = x) 1\{f^*(x) = k\}$$

$$= 1 - \max_{k \in \mathcal{C}} \mathbb{P}(Y = k \mid X = x).$$

- 3. The Bayes classifier assigns x_0 to 1 and the Bayes error at x_0 is 0.5.
- 4. The expected error rate of \hat{f} at x_0 is

$$\mathbb{E}\Big[1\{\hat{f}(X) \neq Y\} \mid X = x_0\Big] = \mathbb{P}(\hat{f}(X) \neq Y \mid X = x_0)$$

$$= 1 - \mathbb{P}(\hat{f}(X) = Y \mid X = x_0)$$

$$= 1 - \sum_{c \in \mathcal{C}} \mathbb{P}(Y = c \mid X = x_0) \mathbb{P}(\hat{f}(X) = c \mid X = x_0)$$

$$= 1 - \sum_{c \in \mathcal{C}} \mathbb{P}(Y = c \mid X = x_0) \mathbb{P}(\hat{f}(X) = c)$$

$$= 1 - \frac{1}{3} \sum_{c \in \mathcal{C}} \mathbb{P}(Y = c \mid X = x_0)$$

$$= 1 - \frac{1}{3} [0.5 + 0.3 + 0.2] = \frac{2}{3},$$

which is greater than the Bayes error.

• Problem 3 (7 pts)

We will prove that a random guessing classifier for binary classification has the area under the curve (AUC) equal to 1/2. Suppose $Y \in \{0,1\}$ with Y = 1 meaning true, and false otherwise. Consider the following random guessing classifier at any X = x

$$\hat{f}(x) = \begin{cases} 1, & \text{with prob. equal to } p \\ 0, & \text{with prob. equal to } 1 - p \end{cases}$$

with any p chosen from [0,1].

- 1. (3 pts) Prove that the expected AUC of this random classifier with p varying within [0,1] is 1/2. [Hint: the expected False Positive Rate is $\mathbb{P}(\hat{f}(X) = 1 \mid Y = 0)$]
- 2. (2 pts) Let $\eta(x) := \mathbb{P}(Y = 1 \mid X = x)$ for any x. Write the expected error rate of \hat{f} at X = x in terms of $\eta(x)$ and p.
- 3. (2 pts) If you have the flexibility of choosing p in your expression of part 2, which choice minimizes the expected error rate of \hat{f} at X = x? Is the resulting classifier equivalent to the Bayes classifier? State your explanation.

SOLUTION:

1. The expected FPR is

$$\mathbb{P}(\hat{f}(X) = 1 \mid Y = 0) = \mathbb{P}(\hat{f}(X) = 1) = p$$

while the expected FNR is

$$\mathbb{P}(\hat{f}(X) = 0 \mid Y = 1) = \mathbb{P}(\hat{f}(X) = 0) = 1 - p.$$

Note that the expected TPR is $1 - \mathbb{P}(\hat{f}(X) = 0 \mid Y = 1) = p$. Therefore, the AUC of \hat{f} is

$$\int_0^1 p dp = \frac{1}{2}.$$

2. The expected error rate at X = x equals to

$$\begin{split} & \mathbb{P}(Y \neq \hat{f}(x) \mid X = x) \\ & = \mathbb{P}(Y = 1, \hat{f}(x) = 0 \mid X = x) + \mathbb{P}(Y = 0, \hat{f}(x) = 1 \mid X = x) \\ & = \mathbb{P}(Y = 1 \mid X = x) \mathbb{P}(\hat{f}(x) = 0 \mid X = x) + \mathbb{P}(Y = 0 \mid X = x) \mathbb{P}(\hat{f}(x) = 1 \mid X = x) \\ & = \mathbb{P}(Y = 1 \mid X = x) \mathbb{P}(\hat{f}(x) = 0) + \mathbb{P}(Y = 0 \mid X = x) \mathbb{P}(\hat{f}(x) = 1) \\ & = \eta(x)(1 - p) + (1 - \eta(x))p. \end{split}$$

3. We choose

$$p = \begin{cases} 1 & \text{if } \eta(x) \ge 1/2\\ 0 & \text{otherwise} \end{cases}$$

The resulting classifier only equals to the Bayes classifier at this point but not in general. Indeed, to match the Bayes classifier for all x, we need to allow p = p(x) being a function that depends on x.

• Problem 4 (14 pts)

This question uses the variables dis (the weighted mean of distances to five Boston employment centers) and nox (nitrogen oxides concentration in parts per 10 million) from the Boston data in the library MASS. We will treat dis as the predictor and nox as the response. For drawing the fitted lines below, let's use the following code to generate a grid of dis

dislims<-range(Boston\$dis)
dis.grid<-seq(from=dislims[1],to=dislims[2],length.out=100)</pre>

- 1. (2 pt) Use the poly() function to fit a polynomial regression with degree equal to 10 to predict nox using dis. Plot the data, the polynomial fit and its 95% confidence band. Comment on the width of the confidence band.
- 2. (2 pts) Plot the polynomial fits for a range of different polynomial degrees, from {1, 3, 5, 7, 10}, and report the associated residual sum of squares (RSS). What's your finding and how to explain it?
- 3. (2 pts) Perform 10-fold cross-validation to select the best degree of the polynomial from {1, 3, 5, 7, 10}.
- 4. (2 pts) Use the bs() function to fit a regression spline with 7 degrees of freedom to predict nox using dis. Specify how the knots are chosen and plot the resulting fit and its 95% confidence band. Comment on the width of confidence band.
- 5. (2 pts) Use the ns() function to fit a natural regression spline with the same knots specified in the previous part. (You may find the nox function and Boundary.knots useful). Plot the resulting fit and its 95% confidence band. Comparing to the previous plot, comment on the difference.
- 6. (2 pts) Now fit natural cubic splines for a range of degrees of freedom, from {5, 10, 15, 20} (specified via df in ns() function), and plot the resulting fits and report the resulting RSS. Comment on what you observe.
- 7. (2 pts) Perform 10-fold cross-validation to select the best degrees of freedom for a natural regression spline from {5, 10, 15, 20}.

• Problem 5 (18 pts)

You will implement different local regression approaches in this problem. Let's set the seed to set.seed(20231101). For n = 300 training data, let's generate (x_i, y_i) for $1 \le i \le n$ from

$$x_i \stackrel{i.i.d.}{\sim} N(3,1), \qquad \epsilon_i \stackrel{i.i.d.}{\sim} N(0,1),$$

 $y_i = 0.5 + 0.1x_i + 0.2x_i^2 + \epsilon_i.$

Let's generate another m = 300 test data points by the same generating mechanism. Set a grid of the training data as $x_grid=seq(0,6,by=0.02)$

1. (5 pts) Implement the k-nearest-neighbor, that is, for any given x_0 , use the averaged responses of its k nearest neighbors to predict, namely,

$$\hat{y}_1(x_0) = \frac{1}{k} \sum_{i \in N_k(x_0)} y_i$$

with $N_k(x_0)$ denoting the index set of k nearest neighbors of x_0 among $\{x_1, \ldots, x_n\}$. Use your function to predict at each value of $\mathbf{x_grid}$. Plot the training data points and add on the lines of your fitted values for $k \in \{5, 20, 50, 100\}$. Make sure you label each line correctly in your plot. Comment on your findings for different choices of k.

2. (5 pts) Repeat step 1 for the weighted k-nn, that is, for any given x_0 , predict by

$$\hat{y}_2(x_0) = \sum_{i \in N_k(x_0)} w(x_0, x_i) y_i.$$

Here we choose the so-called *tricubic weights* as

$$w(x_0, x_i) = \frac{\widetilde{w}(x_0, x_i)}{\sum_{i \in N_k(x_0)} \widetilde{w}(x_0, x_i)}, \quad \forall \ i \in N_k(x_0)$$

$$(0.2)$$

with

$$\widetilde{w}(x_0, x_i) = \left(1 - \left(\frac{|x_i - x_0|}{\max_{i \in N_k(x_0)} |x_i - x_0|}\right)^3\right)^3, \tag{0.3}$$

3. (5 pts) Repeat step 1 for the weighted local linear regression, that is, for any given x_0 ,

$$\hat{y}_3(x_0) = \hat{\beta}_0 + \hat{\beta}_1 x_0$$

with

$$(\hat{\beta}_0, \hat{\beta}_1) = \underset{\beta_0, \beta_1}{\operatorname{argmin}} \sum_{i \in N_k(x_0)} w(x_0, x_i) (y_i - \beta_0 - \beta_1 x_i)^2.$$

Here $w(x_0, x_i)$ is the same as in (0.2) - (0.3). For implementing the above weighted least squares, you might use the 1m function by specifying the weights argument.

4. (3 pts) Use the test data to choose the best one among the above three procedures for $k \in \{5, 20, 50, 100\}$.