### STA 314: Statistical Methods for Machine Learning I

Lecture 2 - Linear regression

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#### Review

 Supervised learning is about to estimate (learn) f under the generating mechanism

$$Y = f(X) + \epsilon$$

- ullet For a given estimate  $\hat{f}$  of f, we have learned
  - how to evaluate it.
  - ▶ and its expected MSE follows the bias-variance decomposition
- Different  $\hat{f}$ 's are different algorithms/methods/predictors:
  - parametric methods
  - non-parametric methods

### Linear regression

Let  $Y \in \mathbb{R}$  be the outcome and  $X \in \mathbb{R}^p$  be the (random) vector of p features.

The linear model assumes

$$Y = f(x) + \epsilon$$
  
=  $\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$  (linearity of  $f$ )

where:

- $\beta_0, \beta_1, \dots, \beta_p$  are unknown constants.
  - $\beta_0$  is called the **intercept**
  - ▶  $\beta_j$ , for  $1 \le j \le p$ , are the **coefficients** or **parameters** of the *p* features
- $\epsilon$  is the error term satisfying  $\mathbb{E}[\epsilon] = 0$ .

### Linear predictor under the linear regression model

Given some estimates  $\hat{\beta}_j$  of  $\beta_j$  for  $0 \le j \le p$ , we predict the response at any  $X = \mathbf{x}$  by the linear predictor

$$\hat{f}(\mathbf{x}) = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p.$$

Question: how to choose  $\hat{\beta}_0, \ldots, \hat{\beta}_p$ ?

### Ordinary Least Squares approach (OLS)

Recall that we want to find a function  $g: \mathcal{X} \to \mathcal{Y}$  by

$$\min_{g} \mathbb{E}\left[\left(Y - g(X)\right)^{2}\right].$$

Under linear model, it suffices to find g by

$$\min_{\alpha_0,\dots,\alpha_p} \mathbb{E}\left[\left(Y - \alpha_0 - \alpha_1 X_1 - \dots - \alpha_p X_p\right)^2\right]$$

In the **model fitting step**, we use the training data to approximate the above expectation (w.r.t. X and Y).

Specifically, given  $\mathcal{D}^{train}$ , we choose

$$(\hat{\beta}_0, \dots, \hat{\beta}_p) = \underset{\alpha_0, \dots, \alpha_p}{\operatorname{argmin}} \frac{1}{n} \sum_{i=1}^n (y_i - \alpha_0 - \alpha_1 x_{i1} - \dots - \alpha_p x_{ip})^2.$$

### Ordinary Least Squares approach (OLS)

Using the matrix notation,

• 
$$\hat{\boldsymbol{\beta}} = [\hat{\beta}_0, \dots, \hat{\beta}_p]^{\top} \in \mathbb{R}^{p+1}, \ \boldsymbol{\beta} = [\beta_0, \dots, \beta_p]^{\top} \in \mathbb{R}^{p+1},$$
  
 $\boldsymbol{\alpha} = [\alpha_0, \dots, \alpha_p]^{\top} \in \mathbb{R}^{p+1},$ 

• 
$$\mathbf{y} = [y_1, \dots, y_n]^{\top} \in \mathbb{R}^n$$
,  $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_n]^{\top} \in \mathbb{R}^{n \times (p+1)}$  with  $\mathbf{x}_i = [1, x_{i1}, \dots, x_{ip}]^{\top} \in \mathbb{R}^{p+1}$  for  $1 \le i \le n$ .

the OLS estimator of  $\beta$  is defined as

$$\hat{\boldsymbol{\beta}} = \underset{\boldsymbol{\alpha} \in \mathbb{R}^{p+1}}{\operatorname{argmin}} \ \frac{1}{n} \sum_{i=1}^{n} \left( y_i - \mathbf{x}_i^{\top} \boldsymbol{\alpha} \right)^2$$
$$= \underset{\boldsymbol{\alpha} \in \mathbb{R}^{p+1}}{\operatorname{argmin}} \ \frac{1}{n} ||\mathbf{y} - \mathbf{X} \boldsymbol{\alpha}||_2^2.$$

#### Comments

 The idea of estimating f by minimizing the training MSE can be applied to (almost) all supervised learning problems. Specifically,

$$\hat{f} = \underset{g \in \mathcal{F}}{\operatorname{argmin}} \frac{1}{n} \sum_{i=1}^{n} L(y_i, g(\mathbf{x}_i))$$
 (1)

where  $L(\cdot, \cdot)$  is a loss function and  $\mathcal{F}$  is a class of choices of g.

• The OLS approach corresponds to  $L(a, b) = (a - b)^2$  and

$$\mathcal{F} = \left\{ \mathbf{x} \mapsto \mathbf{x}^{\top} \boldsymbol{\beta} : \boldsymbol{\beta} \in \mathbb{R}^{p+1} \right\}.$$

 In general, the difficulty of solving (1) varies across problems. But, the OLS approach admits a closed-form solution:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y}$$

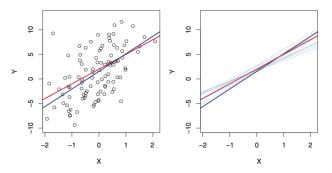
whenever  $\mathbf{X} \in \mathbb{R}^{n \times (p+1)}$  has full column rank.

#### What is random and what is not random?

- In Statistics, we use capitalized letters for generic random variables (e.g. X and Y).
- The parameters such as  $\beta_1, \dots, \beta_p$  or the function  $f : \mathcal{X} \to \mathcal{Y}$  are treated as deterministic (non-random). Of course, being Bayesian is an exception.
- The data points  $(\mathbf{x}_i, y_i)$  for  $1 \le i \le n$  are actual values, observed in practice. They can be thought as the <u>realizations of random variables</u>  $(X_i, Y_i)$  for  $1 \le i \le n$ .
- When we talk about estimators (e.g. the OLS estimator) which, by definition, are functions of  $(X_i, Y_i)$ , hence are random.
- Nevertheless, we will not distinguish between  $(\mathbf{x}_i, y_i)$  and  $(X_i, Y_i)$  throughout the lecture, but you should have in mind that the training data  $(\mathbf{x}_i, y_i)$  are random realizations.

## The randomness in $\hat{\beta}_0$ and $\hat{\beta}_1$

We cannot hope  $\hat{\beta}_0 = \beta_0$  and  $\hat{\beta}_1 = \beta_1$ , because they depend on the observed data which is random.



Left: The **red line** represents the true model f(X) = 2 + 3X. The **blue line** is the OLS fit based on the observed data.

Right: The light blue lines represent 10 OLS fits, each one computed on the basis of a different training dataset.

The fitted OLS lines are different, but their average is close to the true regression line.

### Some important considerations

- Estimation of  $\beta$ :
  - ▶ How close is the point estimation  $\hat{\beta}$  to  $\beta$ ?
  - (Inference) Can we provide confidence interval / conduct hypothesis testing of  $\beta$ ?
- Prediction of Y at  $X = \mathbf{x}$ :
  - ▶ How accurate is the point prediction  $\hat{f}(\mathbf{x}) = \mathbf{x}^{\top} \hat{\boldsymbol{\beta}}$ ?
  - ▶ (Inference) Can we further provide confidence interval of *Y*?
- Variable (Model) selection:
  - ▶ Do all the predictors help to explain *Y*, or is there only a subset of the predictors useful?
  - Later in Lecture 3

# Property of $\hat{\boldsymbol{\beta}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y}$

Take the *design matrix*  $\mathbf{X} \in \mathbb{R}^{n \times (p+1)}$  to be deterministic with full column rank. Assume  $\epsilon_1, \ldots, \epsilon_n$  are i.i.d. realizations of  $\epsilon$ .

- Unbiasedness:  $\mathbb{E}[\hat{\beta}] = \beta$
- The covariance matrix of  $\hat{\beta}$  is:

$$\operatorname{Cov}(\hat{\boldsymbol{\beta}}) = \sigma^2(\mathbf{X}^{\top}\mathbf{X})^{-1}$$

ullet The above two properties imply the  $\ell_2$  estimation error

$$\mathbb{E}\left[\left\|\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}\right\|_2^2\right] = \sigma^2 \mathrm{trace}\left[\left(\mathbf{X}^{\top}\mathbf{X}\right)^{-1}\right]$$

When  $\mathbf{X}^{\mathsf{T}}\mathbf{X} = n \mathbf{I}_{p+1}$ ,

$$\mathbb{E}\left[\|\hat{\boldsymbol{\beta}}-\boldsymbol{\beta}\|_{2}^{2}\right]=\frac{\sigma^{2}(p+1)}{n}.$$

The MSE of estimating  $\beta$  increases as p gets larger.