

Review of a few Probability facts and linear regressions

Xin Bing

Department of Statistical Sciences
University of Toronto

Mathematical notations

- Vector norm: for a vector $v \in \mathbb{R}^d$, its ℓ_p norm, for $0 \leq p \leq \infty$ is defined as

$$\|v\|_p = \left(\sum_{j=1}^d |v_j|^p \right)^{1/p}.$$

We mainly use $\|v\|_1$ and $\|v\|_2$.

- Inner-product between vectors $v_1, v_2 \in \mathbb{R}^d$:

$$v_1^T v_2 = \sum_{j=1}^d v_{1j} v_{2j}.$$

- For any square matrix $M \in \mathbb{R}^{d \times d}$, the trace of M is defined as

$$\text{Tr}(M) = \sum_{j=1}^d M_{jj}$$

In particular, for any vectors $v_1, v_2 \in \mathbb{R}^d$,

$$v_1^\top v_2 = \text{Tr}(v_1 v_2^\top).$$

Let X and Y be two random variables.

- $$\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2.$$

- More generally, for any function f ,

$$\text{Var}(f(X)) = \mathbb{E}[(f(X) - \mathbb{E}[f(X)])^2] = \mathbb{E}[(f(X))^2] - (\mathbb{E}[f(X)])^2.$$

- X is said to be uncorrelated with Y if

$$\text{Cov}(X, Y) = 0.$$

In particular, the fact that X is independent of Y implies that $\text{Cov}(X, Y) = 0$.

- For any constants a, b ,

$$\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab\text{Cov}(X, Y).$$

In particular, if X is uncorrelated with Y , then

$$\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y).$$

- For any function f and g , if X is independent of Y , then

$$\mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)]\mathbb{E}[g(Y)],$$

and

$$\mathbb{E}[f(X) | Y] = \mathbb{E}[f(X)].$$

- For any function h ,

$$\begin{aligned}\mathbb{E}[h(X, Y)] &= \mathbb{E}_X [\mathbb{E}_{Y|X}[h(X, Y) | X]] \\ &= \mathbb{E}_Y [\mathbb{E}_{X|Y}[h(X, Y) | Y]]\end{aligned}$$

where \mathbb{E}_X is the expectation w.r.t. the randomness of X whereas $\mathbb{E}_{Y|X}$ is w.r.t. the randomness of $Y | X$.

Simple Linear Regression

- We assume a model

$$Y = \beta_0 + \beta_1 X + \epsilon,$$

where β_0 and β_1 are two unknown constants that represent the **intercept** and **slope**, also known as **coefficients** or **parameters**, and ϵ is the error term.

- Given some estimates $\hat{\beta}_0$ and $\hat{\beta}_1$ for the model coefficients, we predict response at $X = x$ as

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x.$$

Least Square Estimates

- Training data: $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$.
- Least square estimates $\hat{\beta}_0$ and $\hat{\beta}_1$ are minimizers of RSS , given by

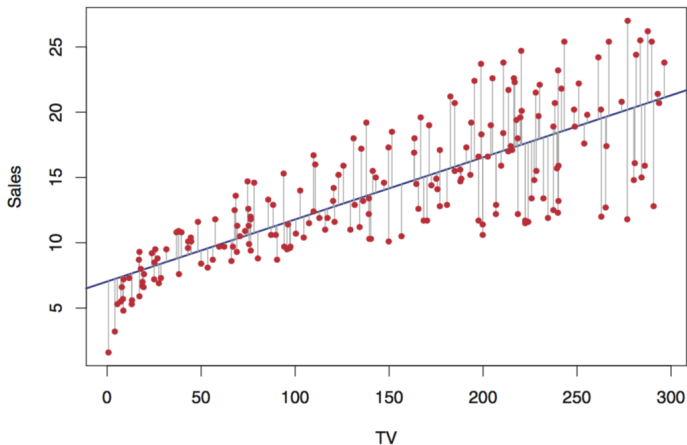
$$(\hat{\beta}_0, \hat{\beta}_1) = \underset{\beta_0, \beta_1}{\operatorname{argmin}} \frac{1}{n} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2.$$

They have the following closed-form solution

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}, \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x},$$

where $\bar{y} = n^{-1} \sum_{i=1}^n y_i$ and $\bar{x} = n^{-1} \sum_{i=1}^n x_i$ are the sample means.

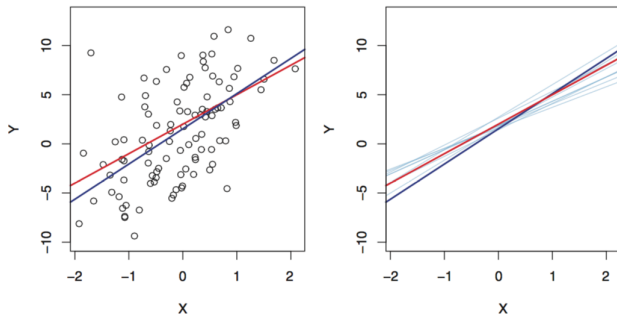
Advertising Data



Each grey line segment represents an error, and the fit makes a compromise by averaging their squares. A linear fit captures the essence of the relationship, although it is somewhat deficient in the left of the plot.

Understand the randomness in $\hat{\beta}_0$ and $\hat{\beta}_1$

We cannot hope $\hat{\beta}_0 = \beta_0$ and $\hat{\beta}_1 = \beta_1$, because they depend on the observed data which is random.



Left: The red line represents the true relationship, $f(X) = 2 + 3X$, which is known as the population regression line. The blue line is the least squares fit based on the observed data.

Right: The light blue lines represent ten least squares fits. Each one is computed on the basis of a different training set.

The fitted least squares lines are different, but their average is quite close to the true regression line.

Derivation of the OLS formula

Recall that

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^{p+1}} \frac{1}{n} \|\mathbf{y} - \mathbf{X}\beta\|_2^2.$$

Taking the derivative with respect to β and setting it equal to zero yield

$$-\frac{2}{n} \mathbf{X}^\top (\mathbf{y} - \mathbf{X}\beta) = 0.$$

The solution $\hat{\beta}$ has to satisfy the above equation.

When \mathbf{X} has full column rank such that $\mathbf{X}^\top \mathbf{X}$ is invertible, there exists a unique solution, i.e.

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}.$$