## Homework 3 (Nov. 1st)

Deadline: Sunday, November 19th, at 11:59pm.
Submission: There are five questions in this homework. You need to submit separate PDF files to each question via Crowdmark. For Questions $4 \& 5$, your submission should also contain the R code and R outputs. You can produce the PDFs however you like (e.g. LATEX, Microsoft Word, scanner), as long as they are legible.

Neatness Point: You will be deducted one point if we have a hard time reading your solutions or understanding the structure of your code.

Late Submission: $10 \%$ of the total possible marks will be deducted for each day late, up to a maximum of 3 days. After that, no submissions will be accepted.

## - Problem 1 (4 pts)

It was mentioned in the lecture that a cubic regression spline with two knots at $\xi_{1}<\xi_{2}$ can be represented as

$$
f(x)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\beta_{3} x^{3}+\beta_{4}\left(x-\xi_{1}\right)_{+}^{3}+\beta_{5}\left(x-\xi_{2}\right)_{+}^{3}
$$

where

$$
\left(x-\xi_{k}\right)_{+}^{3}=\left\{\begin{array}{ll}
\left(x-\xi_{k}\right)^{3} & \text { if } x>\xi_{k}  \tag{0.1}\\
0 & \text { otherwise }
\end{array}, \quad \forall k \in\{1,2\}\right.
$$

We now verify $f(x)$ is indeed a cubic regression spline by proving:

1. ( $\mathbf{1} \mathbf{~ p t ) ~} f(x)$ can be written as three cubic polynomials for $x \leq \xi_{1}, \xi_{1}<x \leq \xi_{2}$ and $x>\xi_{2}$.
2. ( $\mathbf{1} \mathbf{~ p t ) ~ t h e ~ t h r e e ~ p o l y n o m i a l s ~ h a v e ~ t h e ~ s a m e ~ v a l u e ~ a t ~ e a c h ~ k n o t . ~}$
3. ( $\mathbf{1} \mathbf{p t ) ~ t h e ~ t h r e e ~ p o l y n o m i a l s ~ h a v e ~ t h e ~ s a m e ~ f i r s t ~ o r d e r ~ d e r i v a t i v e ~ a t ~ e a c h ~ k n o t . ~}$
4. ( $\mathbf{1} \mathbf{~ p t ) ~ t h e ~ t h r e e ~ p o l y n o m i a l s ~ h a v e ~ t h e ~ s a m e ~ s e c o n d ~ o r d e r ~ d e r i v a t i v e ~ a t ~ e a c h ~ k n o t . ~}$

## SOLUTION:

1. According to the definition of $f(x)$, we have

$$
f(x)= \begin{cases}f_{1}(x)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\beta_{3} x^{3} & \text { if } x \leq \xi_{1} \\ f_{2}(x)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\beta_{3} x^{3}+\beta_{4}\left(x-\xi_{1}\right)^{3} & \text { if } \xi_{1}<x \leq \xi_{2} \\ f_{3}(x)=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\beta_{3} x^{3}+\beta_{4}\left(x-\xi_{1}\right)^{3}+\beta_{5}\left(x-\xi_{2}\right)^{3} & \text { if } x>\xi_{2}\end{cases}
$$

as desired.
2. This follows immediately from the previous part by letting $x=\xi_{1}$ and $x=\xi_{2}$.
3. Since

$$
\begin{aligned}
& \frac{d f_{1}(x)}{d x}=\beta_{1}+2 \beta_{2} x+3 \beta_{3} x^{2}, \\
& \frac{d f_{2}(x)}{d x}=\beta_{1} x+2 \beta_{2} x+3 \beta_{3} x^{2}+3 \beta_{4}\left(x-\xi_{1}\right)^{2}, \\
& \frac{d f_{3}(x)}{d x}=\beta_{1} x+2 \beta_{2} x+3 \beta_{3} x^{2}+3 \beta_{4}\left(x-\xi_{1}\right)^{2}+3 \beta_{5}\left(x-\xi_{2}\right)^{2},
\end{aligned}
$$

$f_{1}^{\prime}\left(\xi_{1}\right)=f_{2}^{\prime}\left(\xi_{1}\right)$ and $f_{2}^{\prime}\left(\xi_{2}\right)=f_{3}^{\prime}\left(\xi_{2}\right)$ follow by taking $x=\xi_{1}$ and $x=\xi_{2}$.
4. Since

$$
\begin{aligned}
& \frac{d^{2} f_{1}(x)}{d x^{2}}=2 \beta_{2}+6 \beta_{3} x \\
& \frac{d^{2} f_{2}(x)}{d x^{2}}=2 \beta_{2}+6 \beta_{3} x+6 \beta_{4}\left(x-\xi_{1}\right) \\
& \frac{d^{2} f_{2}(x)}{d x^{2}}=2 \beta_{2}+6 \beta_{3} x+6 \beta_{4}\left(x-\xi_{1}\right)+6 \beta_{5}\left(x-\xi_{2}\right)
\end{aligned}
$$

taking $x=\xi_{1}$ and $x=\xi_{2}$ leads to $f_{1}^{\prime \prime}\left(\xi_{1}\right)=f_{2}^{\prime \prime}\left(\xi_{1}\right)$ and $f_{2}^{\prime \prime}\left(\xi_{2}\right)=f_{3}^{\prime \prime}\left(\xi_{2}\right)$.

## - Problem 2 ( 7 pts)

Consider the classification problem with the label of $Y$ belong to $\mathcal{C}:=\{1,2, \ldots, K\}$ and any realization $x$ of $X \in \mathbb{R}^{p}$. The Bayes classifier $f^{*}$ at $X=x$ is defined as

$$
f^{*}(x):=\underset{f(x) \in \mathcal{C}}{\operatorname{argmin}} \mathbb{E}[1\{Y \neq f(x)\} \mid X=x] .
$$

In the class, we have shown that

$$
f^{*}(x)=\underset{f(x) \in \mathcal{C}}{\operatorname{argmax}} \mathbb{P}(Y=f(x) \mid X=x) .
$$

1. (2 pts) Prove that

$$
f^{*}(x)=\underset{k \in \mathcal{C}}{\operatorname{argmax}} \mathbb{P}(Y=k \mid X=x) .
$$

2. (1 pt) Prove that the Bayes error at $X=x$ equals to

$$
1-\max _{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x)
$$

3. ( $\mathbf{2} \mathbf{~ p t s}$ ) Consider that $K=3$. For a fixed $x_{0}$, assume that

$$
\begin{aligned}
& \mathbb{P}\left(Y=1 \mid X=x_{0}\right)=0.5 \\
& \mathbb{P}\left(Y=2 \mid X=x_{0}\right)=0.3 \\
& \mathbb{P}\left(Y=3 \mid X=x_{0}\right)=0.2
\end{aligned}
$$

State the Bayes classifier at $X=x_{0}$ and compute its error at $X=x_{0}$.
4. ( $\mathbf{2} \mathbf{~ p t s}$ ) Consider a naive classifier $\hat{f}$, called random guessing, which randomly picks one label from $\mathcal{C}=\{1,2,3\}$ with equal probability. Compute its expected error rate at $X=x_{0}$ and compare it with the Bayes error at $X=x_{0}$ in part 3 .

## SOLUTION:

1. Proof. We have

$$
\begin{aligned}
\mathbb{P}(Y=f(x) \mid X=x) & =\sum_{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x) \mathbb{P}(f(x)=k \mid X=x) \\
& =\sum_{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x) 1\{f(x)=k\} \\
& \leq \max _{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x) .
\end{aligned}
$$

Moreover, the maximal value is achieved when $f(x)=k^{*}$ for

$$
k^{*}=\underset{k \in \mathcal{C}}{\operatorname{argmax}} \mathbb{P}(Y=k \mid X=x) .
$$

This proves the first claim.
2. Proof. Regarding the second claim, by similar reasoning,

$$
\begin{aligned}
\mathbb{P}\left(Y \neq f^{*}(x) \mid Y=x\right) & =1-\mathbb{P}\left(Y=f^{*}(x) \mid Y=x\right) \\
& =1-\sum_{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x) 1\left\{f^{*}(x)=k\right\} \\
& =1-\max _{k \in \mathcal{C}} \mathbb{P}(Y=k \mid X=x) .
\end{aligned}
$$

3. The Bayes classifier assigns $x_{0}$ to 1 and the Bayes error at $x_{0}$ is 0.5 .
4. The expected error rate of $\hat{f}$ at $x_{0}$ is

$$
\begin{aligned}
\mathbb{E}\left[1\{\hat{f}(X) \neq Y\} \mid X=x_{0}\right] & =\mathbb{P}\left(\hat{f}(X) \neq Y \mid X=x_{0}\right) \\
& =1-\mathbb{P}\left(\hat{f}(X)=Y \mid X=x_{0}\right) \\
& =1-\sum_{c \in \mathcal{C}} \mathbb{P}\left(Y=c \mid X=x_{0}\right) \mathbb{P}\left(\hat{f}(X)=c \mid X=x_{0}\right) \\
& =1-\sum_{c \in \mathcal{C}} \mathbb{P}\left(Y=c \mid X=x_{0}\right) \mathbb{P}(\hat{f}(X)=c) \\
& =1-\frac{1}{3} \sum_{c \in \mathcal{C}} \mathbb{P}\left(Y=c \mid X=x_{0}\right) \\
& =1-\frac{1}{3}[0.5+0.3+0.2]=\frac{2}{3}
\end{aligned}
$$

which is greater than the Bayes error.

## - Problem 3 ( 7 pts)

We will prove that a random guessing classifier for binary classification has the area under the curve (AUC) equal to $1 / 2$. Suppose $Y \in\{0,1\}$ with $Y=1$ meaning true, and false otherwise. Consider the following random guessing classifier at any $X=x$

$$
\hat{f}(x)=\left\{\begin{array}{ll}
1, & \text { with prob. equal to } p \\
0, & \text { with prob. equal to } 1-p
\end{array} .\right.
$$

1. ( $\mathbf{3} \mathbf{p t s}$ ) Prove that the expected AUC of $\hat{f}$ is $1 / 2$. [Hint: the expected False Positive Rate is $\mathbb{P}(\hat{f}(X)=1 \mid Y=0)$ ]
2. (2 pts) Let $\eta(x):=\mathbb{P}(Y=1 \mid X=x)$ for any $x$. Write the expected error rate of $\hat{f}$ at $X=x$ in terms of $\eta(x)$ and $p$.
3. ( $\mathbf{2} \mathbf{~ p t s}$ ) If you have the flexibility of choosing $p$ in your expression of part 2 , which choice minimizes the expected error rate of $\hat{f}$ at $X=x$ ? Is the resulting classifier equivalent to the Bayes classifier? State your explanation.

## SOLUTION:

1. The expected FPR is

$$
\mathbb{P}(\hat{f}(X)=1 \mid Y=0)=\mathbb{P}(\hat{f}(X)=1)=p
$$

while the expected FNR is

$$
\mathbb{P}(\hat{f}(X)=0 \mid Y=1)=\mathbb{P}(\hat{f}(X)=0)=1-p
$$

Note that the expected TPR is $1-\mathbb{P}(\hat{f}(X)=0 \mid Y=1)=p$. Therefore, the AUC of $\hat{f}$ is

$$
\int_{0}^{1} p d p=\frac{1}{2} p^{2}=\frac{1}{2} .
$$

2. The expected error rate at $X=x$ equals to

$$
\begin{aligned}
& \mathbb{P}(Y \neq \hat{f}(x) \mid X=x) \\
& =\mathbb{P}(Y=1, \hat{f}(x)=0 \mid X=x)+\mathbb{P}(Y=0, \hat{f}(x)=1 \mid X=x) \\
& =\mathbb{P}(Y=1 \mid X=x) \mathbb{P}(\hat{f}(x)=0 \mid X=x)+\mathbb{P}(Y=0 \mid X=x) \mathbb{P}(\hat{f}(x)=1 \mid X=x) \\
& =\mathbb{P}(Y=1 \mid X=x) \mathbb{P}(\hat{f}(x)=0)+\mathbb{P}(Y=0 \mid X=x) \mathbb{P}(\hat{f}(x)=1) \\
& =\eta(x)(1-p)+(1-\eta(x)) p .
\end{aligned}
$$

3. We choose

$$
p= \begin{cases}1 & \text { if } \eta(x) \geq 1 / 2 \\ 0 & \text { otherwise }\end{cases}
$$

The resulting classifier only equals to the Bayes classifier at this point but not in general.

## - Problem 4 ( 14 pts)

This question uses the variables dis (the weighted mean of distances to five Boston employment centers) and nox (nitrogen oxides concentration in parts per 10 million) from the Boston data in the library MASS. We will treat dis as the predictor and nox as the response. For drawing the fitted lines below, let's use the following code to generate a grid of dis
dislims<-range (Boston\$dis)
dis.grid<-seq(from=dislims[1], to=dislims [2],length.out=100)

1. ( $\mathbf{2} \mathbf{~ p t}$ ) Use the poly () function to fit a polynomial regression with degree equal to 10 to predict nox using dis. Plot the data, the polynomial fit and its $95 \%$ confidence band. Comment on the width of the confidence band.
2. ( $\mathbf{2} \mathbf{~ p t s ) ~ P l o t ~ t h e ~ p o l y n o m i a l ~ f i t s ~ f o r ~ a ~ r a n g e ~ o f ~ d i f f e r e n t ~ p o l y n o m i a l ~ d e g r e e s , ~ f r o m ~}\{1,3$, $5,7,10\}$, and report the associated residual sum of squares (RSS). What's your finding and how to explain it?
3. (2 pts) Perform 10 -fold cross-validation to select the best degree of the polynomial from $\{1,3,5,7,10\}$.
4. (2 pts) Use the bs() function to fit a regression spline with 7 degrees of freedom to predict nox using dis. Specify how the knots are chosen and plot the resulting fit and its $95 \%$ confidence band. Comment on the width of confidence band.
5. (2 pts) Use the ns () function to fit a natural regression spline with the same knots specified in the previous part. (You may find the nox function and Boundary.knots useful). Plot the resulting fit and its $95 \%$ confidence band. Comparing to the previous plot, comment on the difference.
6. ( $\mathbf{2} \mathbf{~ p t s ) ~ N o w ~ f i t ~ n a t u r a l ~ c u b i c ~ s p l i n e s ~ f o r ~ a ~ r a n g e ~ o f ~ d e g r e e s ~ o f ~ f r e e d o m , ~ f r o m ~ \{ 5 , ~ 1 0 , ~}$ $15,20\}$ (specified via df in $\mathrm{ns}($ ) function), and plot the resulting fits and report the resulting RSS. Comment on what you observe.
7. ( $\mathbf{2} \mathbf{~ p t s}$ ) Perform 10-fold cross-validation to select the best degrees of freedom for a natural regression spline from $\{5,10,15,20\}$.

## - Problem 5 ( 18 pts)

You will implement different local regression approaches in this problem. Let's set the seed to set.seed(20231101). For $n=300$ training data, let's generate ( $x_{i}, y_{i}$ ) for $1 \leq i \leq n$ from

$$
\begin{aligned}
& x_{i} \stackrel{i . i . d .}{\sim} N(3,1), \quad \epsilon_{i} \stackrel{i . i . d .}{\sim} N(0,1), \\
& y_{i}=0.5+0.1 x_{i}+0.2 x_{i}^{2}+\epsilon_{i} .
\end{aligned}
$$

Let's generate another $m=300$ test data points by the same generating mechanism. Set a grid of the training data as $x_{-g r i d=} \operatorname{seq}(0,6, b y=0.02)$

1. ( $\mathbf{5} \mathbf{~ p t s}$ ) Implement the $k$-nearest-neighbor, that is, for any given $x_{0}$, use the averaged responses of its $k$ nearest neighbors to predict, namely,

$$
\hat{y}_{1}\left(x_{0}\right)=\frac{1}{k} \sum_{i \in N_{k}\left(x_{0}\right)} y_{i}
$$

with $N_{k}\left(x_{0}\right)$ denoting the index set of $k$ nearest neighbors of $x_{0}$ among $\left\{x_{1}, \ldots, x_{n}\right\}$.
Use your function to predict at each value of x_grid. Plot the training data points and add on the lines of your fitted values for $k \in\{5,20,50,100\}$. Make sure you label each line correctly in your plot. Comment on your findings for different choices of $k$.
2. ( $\mathbf{5} \mathbf{~ p t s}$ ) Repeat step 1 for the weighted $k$-nn, that is, for any given $x_{0}$, predict by

$$
\hat{y}_{2}\left(x_{0}\right)=\frac{1}{k} \sum_{i \in N_{k}\left(x_{0}\right)} w\left(x_{0}, x_{i}\right) y_{i} .
$$

Here we choose the so-called tricubic weights as

$$
\begin{equation*}
w\left(x_{0}, x_{i}\right)=\frac{\widetilde{w}\left(x_{0}, x_{i}\right)}{\sum_{i \in N_{k}\left(x_{0}\right)} \widetilde{w}\left(x_{0}, x_{i}\right)}, \quad \forall i \in N_{k}\left(x_{0}\right) \tag{0.2}
\end{equation*}
$$

with

$$
\begin{equation*}
\widetilde{w}\left(x_{0}, x_{i}\right)=\left(1-\left(\frac{\left|x_{i}-x_{0}\right|}{\max _{i \in N_{k}\left(x_{0}\right)}\left|x_{i}-x_{0}\right|}\right)^{3}\right)^{3}, \tag{0.3}
\end{equation*}
$$

3. ( $\mathbf{5} \mathbf{~ p t s}$ ) Repeat step 1 for the weighted local linear regression, that is, for any given $x_{0}$,

$$
\hat{y}_{3}\left(x_{0}\right)=\hat{\beta}_{0}+\hat{\beta}_{1} x_{0}
$$

with

$$
\left(\hat{\beta}_{0}, \hat{\beta}_{1}\right)=\underset{\beta_{0}, \beta_{1}}{\operatorname{argmin}} \sum_{i \in N_{k}\left(x_{0}\right)} w\left(x_{0}, x_{i}\right)\left(y_{i}-\beta_{0}-\beta_{1} x_{i}\right)^{2} .
$$

Here $w\left(x_{0}, x_{i}\right)$ is the same as in (0.2) - (0.3). For implementing the above weighted least squares, you might use the lm function by specifying the weights argument.
4. ( $\mathbf{3} \mathbf{~ p t s}$ ) Use the test data to choose the best one among the above three procedures for $k \in\{5,20,50,100\}$.

